

Supplementary Material for “Asymptotics of score test in the generalized β -model for networks”

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One lemma used in the proof of Theorem 2

We say an $n \times n$ matrix $V_n = (v_{ij})$ belongs to a matrix class $\mathcal{L}_n(m, M)$ if V_n is a symmetric nonnegative matrix satisfying

$$v_{ii} = \sum_{j=1; j \neq i}^n v_{ij}; \quad M \geq v_{ij} = v_{ji} \geq m > 0, \quad i \neq j.$$

Yan et al. (2015) proposed a simple matrix $S_n = \text{diag}(1/v_{11}, \dots, 1/v_{nn})$ to approximate the inverse of V_n and obtained the upper bounds on approximation errors below.

Lemma 1. *If $V_n \in \mathcal{L}_n(m, M)$, then for $n \geq 3$, the following holds:*

$$\|W_n := V_n^{-1} - S_n\| \leq \frac{M(nM + (n-2)m)}{2m^3(n-2)(n-1)^2} + \frac{1}{2m(n-1)^2} + \frac{1}{mn(n-1)},$$

where $\|A\| = \max_{i,j} |a_{ij}|$ denotes a matrix norm for a general matrix $A = (a_{ij})$.

Calculation of equation (2)

The variance of $\sum_i c_i \tilde{d}_i^2$ can be divided into two parts:

$$\text{Var}\left(\sum_i c_i \tilde{d}_i^2\right) = \text{Cov}\left(\sum_i c_i \tilde{d}_i^2, \sum_j c_j \tilde{d}_j^2\right) = \sum_i c_i^2 \text{Var}(\tilde{d}_i^2) + 2 \sum_{1 \leq i < j \leq n} c_i c_j \text{Cov}(\tilde{d}_i^2, \tilde{d}_j^2). \quad (1)$$

We will calculate the first part:

$$\text{Var}(\tilde{d}_i^2) = \text{Cov}\left(\left(\sum_{\alpha=1}^n \tilde{a}_{i\alpha}\right)^2, \left(\sum_{\zeta=1}^n \tilde{a}_{i\zeta}\right)^2\right) = \text{Cov}\left(\sum_{\alpha=1}^n \sum_{\beta=1}^n \tilde{a}_{i\alpha} \tilde{a}_{i\beta}, \sum_{\zeta=1}^n \sum_{\eta=1}^n \tilde{a}_{i\zeta} \tilde{a}_{i\eta}\right).$$

Note that the random variables \tilde{a}_{ij} for $1 \leq i < j \leq n$ are mutually independent. There are only two cases for which $\text{Cov}(\tilde{a}_{i\alpha} \tilde{a}_{i\beta}, \tilde{a}_{i\zeta} \tilde{a}_{i\eta})$ is not equal to zero: (1) $\alpha = \beta = \zeta = \eta \neq i$;

(2) $\alpha = \zeta, \beta = \eta$ or $\alpha = \eta, \beta = \zeta$. By calculation, we have

$$\text{Var}(\tilde{d}_i^2) = 2v_{ii}^2 + \sum_{j \neq i} u_{ij}. \quad (2)$$

The second part of (1) can be calculated as follows.

$$\text{Cov}(\tilde{d}_i^2, \tilde{d}_j^2) = \text{Cov}\left(\left(\sum_{\alpha=1}^n \tilde{a}_{i\alpha}\right)^2, \left(\sum_{\zeta=1}^n \tilde{a}_{j\zeta}\right)^2\right) = \text{Cov}\left(\sum_{\alpha=1}^n \sum_{\beta=1}^n \tilde{a}_{i\alpha} \tilde{a}_{i\beta}, \sum_{\zeta=1}^n \sum_{\eta=1}^n \tilde{a}_{j\zeta} \tilde{a}_{j\eta}\right).$$

In the above, the only case for $\text{Cov}(\tilde{a}_{i\alpha} \tilde{a}_{i\beta}, \tilde{a}_{j\zeta} \tilde{a}_{j\eta})$ not equal to 0, is $\alpha = \beta = j$ and $\zeta = \eta = i$. Then

$$\text{Cov}(\tilde{d}_i^2, \tilde{d}_j^2) = E(\tilde{d}_{ij}^4) - (E(\tilde{d}_{ij}^2))^2 = u_{ij}. \quad (3)$$

Combing (2) and (3) into (1), it yields equation (2).

References

- Yan T., Zhao Y., and Qin H. (2015). Asymptotic normality in the maximum entropy models on graphs with an increasing number of parameters. *J. Multivariate Anal.* **133**, 61–76.